It might be asked why we have more senses than one. [Had it been otherwise],... everything would have merged for us into an indistinguishable identity.

Aristotle, De Anima (350 B.C.E)

Chapter 3 Perceptual Grounding

Most of the enormous variability in the world around us is unimportant. Variations in our sensory perceptions are not only tolerated, they generally pass unnoticed. Of course, some distinctions are of paramount importance and learning which are meaningful as opposed to which can be safely ignored is a fundamental problem of cognitive development. This process is a component of *perceptual grounding*, where a perceiver learns to make sense of its sensory inputs. The perspective taken here is that this is a clustering problem, in that each sense must learn to organize its perceptions into meaningful categories. That animals do this so readily belies its complexity. For example, people learn phonetic structures for languages simply by listening to them; the phonemes are somehow extracted and clustered from auditory inputs even though the listener does not know in advance how many unique phonemes are present in the signal.

Contrast this with a standard mathematical approach to clustering, where some knowledge of the clusters, e.g., how many there are or their distributions, must be known a priori in order to derive them. Without knowing these parameters in advance, algorithmic clustering techniques may not be robust (Fraley and Raftery 2002, Kleinberg 2002, Still and Bialek 2004). Assuming that in many circumstances animals cannot know the parameters underlying their perceptual inputs, how then do they learn to organize their sensory perceptions reliably?

This chapter presents an approach to clustering based on observed correlations between different sensory modalities. These cross-modal correlations exist because perceptions are created through physical processes governed by natural laws (Thompson 1917, Richards 1980, Mumford 2004). An event in the world is simultaneously perceived through multiple sensory pathways in a single observer; while each pathway may have a

unique perspective on the event, their perspectives tend to be correlated by regularities in the physical world (Richards and Bobick 1988). We propose here that these correspondences play a primary role in organizing the sensory channels individually. Based on this hypothesis, we develop a new framework for grounding artificial perceptual systems.

Towards this, we will introduce a mathematical model of *slices*, which are topological manifolds that encode dynamic perceptual states and are inspired by surface models of cortical tissue (Dale et al. 1999, Fischl et al. 1999, Citti and Sarti 2003, Ratnanather et al. 2003). Slices partition perceptual spaces into large numbers of small regions (hyperclusters) and then reassemble them to construct clusters corresponding to the actual sensory events being perceived. This reassembly is performed by *cross-modal clustering*, which uses temporal correlations between slices to determine which hyperclusters within a slice correspond to the same sensory events. The cross-modal clustering algorithm does not presume that either the number of clusters in the data or their distributions is known beforehand. We examine the outputs and behavior of this algorithm on simulated datasets, drawn from a variety of mixture distributions, and on real data gathered in computational experiments.

The Simplest Complex Example

As in Chapter 2, we proceed here by first considering an example. We will return to using real datasets towards the end of this chapter, but for the moment, it is helpful to pare down the subject matter to its bare essentials.

Let us consider two hypothetical sensory modes, each of which is capable of sensing the same two events in the world, which we call the *red* and *blue* events. These two modes are illustrated below in Figure 11, where the dots within each mode represent its perceptual inputs and the blue and red ellipses delineate the two events. For example, if a "red" event takes place in the world, each mode would receive sensory input that (probabilistically) falls within its red ellipse. Notice that events within each mode



Figure 11 – Two hypothetical co-occurring perceptual modes. Each mode, unbeknownst to itself, receives inputs generated by a simple, overlapping Gaussian mixture model. To make matters more concrete, we might imagine Mode A is a simple auditory system that hears two different events in the world and Mode B is a simple visual system sees those same two events, which are indicated by the red and

overlap, and they are in fact represented by a mixture of two overlapping Gaussian distributions. We have chosen this example because it is simple – each mode perceives only two events – but it has the added complexity that the events overlap – meaning there is likely to be some ambiguity in interpreting the perceptual inputs.

Keep in mind that while *we* know there are only two events (red and blue) in this hypothetical world, the *modes* themselves do not "know" anything at all about what they can perceive. The colorful ellipses are solely for the reader's benefit; the only thing the modes receive is their raw input data. Our goal then is to learn the perceptual categories in each mode – e.g., to learn that each mode in this example senses these two overlapping events – by exploiting the temporal correlations between them.

Generating Codebooks

We are going to proceed by hyperclustering each perceptual space into a codebook. This simply means that we are going to generate far more clusters than are necessary for representing the actual number of perceptual events in the data. In this case, that would be two, but instead, we will employ a (much) larger number. For the rest of this discussion, we will refer to two different types of clusters:



Figure 12 – Hyperclustering Mode B with the algorithm given below. Mode B is shown hyperclustered on the right. Here, we specified k=30 and the algorithm ended up generating 53 clusters after normalizing their densities.

- 1) *codebook clusters* (or *hyperclusters*) are generated by hyperclustering and are illustrated by the Voronoi regions show in Figure 12 on the right.
- 2) *perceptual clusters* refer to actual sensory events and are outlined with the colored ellipses in Figure 11.

Our goal will be to combine the *codebook* clusters to "assemble" the *perceptual* clusters. We note that while perceptual clustering is quite difficult, for reasons outlined in the previous chapter, hyperclustering is quite easy because there is no notion of perceptual correctness associated with it. Although we must determine how many codebook clusters to generate, we will show this number influences the amount of training data necessary rather than the correctness of the derived perceptual clusters. In other words, this approach is not overly sensitive to the hyperclustering: generating too many hyperclusters simply means learning takes longer, not that the end results are incorrect. Generating too few hyperclusters to happen because of the density normalization described below. It is also sometimes possible to detect that too few clusters have been generated by using cross-modal information, a technique we examine later in this chapter.

To generate the codebooks, we will use a variant of the Generalized Lloyd Algorithm (GLA) (Lloyd 1982). We modify the algorithm to normalize the point densities within the hyperclusters, which otherwise can vary enormously. Many clustering algorithms, including GLA, optimize initially random codebooks by minimizing a strongly Euclidean distance metric between cluster centroids and their members. A cluster with a large

numbers of nearby points may be viewed as equivalent to (from the perspective of the optimization) a cluster with a small number of distant points. It is therefore possible to have substantial variance in the number of points assigned to each codebook cluster. This is problematic because our approach will require that each *perceptual* cluster be represented by multiple *codebook* clusters, from which it is "assembled." The Euclidean bias introduced by the distance metric used for codebook optimization means that "small" perceptual events may be relegated to a single codebook cluster. This would prevent them from ever being detected.

There are many ways one could imagine achieving this density normalization. For example, we could explicitly add inverse cluster size to the minimization calculation performed during codebook refinement. This would leave the number of codebook clusters constant overall but introduce pressure against wide variation in the number of points assigned to each one. Rather than take an approach that preserves the overall number of clusters, we will instead modify the algorithm to recluster codebook regions that have been assigned "too many" points. This benefit of this is that we leave the GLA algorithm intact but now invoke it recursively on subregions where its performance is unsatisfactory. By keeping the basic structure of GLA, many of the mathematic properties of the generated codebooks remain unchanged. The downside of this approach is that the recursive reclustering increases the total number of generated hyperclusters. Thus, the algorithm generates at least as many codebook clusters as we specify and sometimes many more. This increase in codebook size can affect the computational complexity of algorithms operating over these codebooks, which we investigate later in this chapter. We note, however, that adding these additional clusters does not tend to require gathering more training data, an issue raised above. This is because the extra clusters are generated in regions that already have high point densities.

Our hyperclustering algorithm for generating (at least) k codebook regions over dataset $D \subseteq \mathbb{R}^{N}$ is:

1) Let s = |D|/k. This is our goal size for the number of data points per cluster.



Figure 13 - The hyperclusters generated for the data in Mode B, with the data removed. The number identifying each cluster is located at its centroid. Notice how the number of clusters increases in the region corresponding to the overlap of the two Gaussian distributions, where the overlap point density is highest.

- 2) Let $P = \{P_1, P_2, ..., P_k\}, P_i \subset \mathbb{R}^N$ be a Lloyd partitioning of *D* over *k* clusters. This is the output of the Generalized Lloyd Algorithm.
- 3) For each cluster $P_i \in P$:

If $|P_i| > s$ (the cluster has too many points), then recursively partition P_i :

- a. Let $Q = \{P_1, P_2\}, P_i \subset \mathbb{R}^N$ be a Lloyd partitioning of P_i over 2 clusters.
- b. Set $P = (P \cup Q) / P_i$. Add the two new partitions and remove the old one.

End if statement

4) Repeat step 3 until no new partitions are added. Then, return the centroids of the sets in *P* as the final hyperclustering. Empirically, we find that k < |P| < 2k.

The output of this algorithm on the data in Mode B is shown above in Figure 13. Notice how the number of clusters increases in the region corresponding to the overlap of the two Gaussian distributions, which is due to the density normalization. We note that any number of variations on this algorithm are possible. For example, in the reclustering step in (3), we might recursively generate $|P_i|/s$ rather than 2 clusters. We could also modify the goal size *s* to change the degree of density normalization. In any event, we have



Figure 14 – Slices generated for Modes A and B using the hyperclustering algorithm in the previous section. Our goal is to combine the codebook clusters to reconstruct the actual sensory events perceived within the slices.

found that our approach is not particularly sensitive to the precise details of the codebook's generation; we confirm this statement later in this chapter, when we consider hyperclustering other mixture distributions. At present, the most important consideration is that the cluster densities are normalized, which minimizes the Euclidean bias inherent in the centroid optimization performed by the Lloyd algorithm.

Generating Slices

We now introduce a new data structure called *slices* that are constructed using the codebooks defined in the previous section. Figure 14 illustrates slices constructed for Modes A and B from our example above. Slices are topological manifolds that encode dynamic perceptual states and are inspired by surface models of cortical tissue (Citti and Sarti 2003, Ratnanather et al. 2003). They are able to represent both symbolic and numeric data and provide a natural foundation for aggregating and correlating information. Intuitively, a slice is a codebook with a non-Euclidean distance metric defined between its cluster centroids. In other words, distances within each cluster are Euclidean, whereas distances between clusters are not. A topological manifold is simply a manifold "glued" together from Euclidean spaces, and that is exactly what a slice is.

Our goal is to combine the codebook regions to "reconstruct" the larger perceptual regions within a slice. To do this, we will define a non-Euclidean distance metric



Figure 15 – Viewing Hebbian linkages between two different slices. The modes have been vertically stacked here to make the correspondences clearer. The blue lines indicate that two codebook regions temporally co-occur with each other. Note that these connections are weighted based on their strengths, which is not visually represented here, and that these weights are additionally asymmetric between each pair of connected regions.

between codebook regions that reflects how much we think they are part of the same perceptual event. In this metric, codebook regions corresponding to the same perceptual event will be closer together and those corresponding to different events will be further apart. Towards defining this metric, we first collect co-occurrence data between the codebook regions in different modes. We want to know how each codebook region in a mode temporally co-occurs with the codebook regions in other modes.

This data can be easily gathered with the classical sense of Hebbian learning (Hebb 1949), where connections between regions are strengthened as they are simultaneously active. The result of this process is illustrated in Figure 15, where the modes are vertically stacked to make the correspondences clearer. We will exploit the spatial structure of this Hebbian co-occurrence data to define the distance metric within each mode.

Hebbian Projections

In this section, we define the notion of a *Hebbian projection*. These are spatial probability distributions that provide an intuitive way to view co-occurrence relations between different slices. We first give a formal definition and then illustrate the concept visually.

Consider two slices $M_A, M_B \subseteq \mathbb{R}^n$, with associated codebooks $C_A = \{p_1, p_2, ..., p_a\}$ and $C_B = \{q_1, q_2, ..., q_b\}$, where cluster centroids $p_i, q_j \in \mathbb{R}^N$.

For some event x, we define h(x) = # of times event x occurs. Similarly, for events x and y, we define h(x, y) = # of times events x and y co-occur. For example, $h(p_1)$ is the number of times inputs that belong to cluster p_1 were seen during some time period of interest. So, we see that Pr(x | y) = h(p,q)/h(p).

We define the *Hebbian projection* of a codebook cluster $p_i \in C_A$ onto mode M_B :

$$\dot{H}_{A}^{B}(p_{i}) = \left[\Pr(q_{1} \mid p_{i}), \Pr(q_{2} \mid p_{i}), ..., \Pr(q_{b} \mid p_{i})\right]$$
(3.1)

When the modes are clear from context, we will simply refer to the projection by $\vec{H}(p_i)$.

A Hebbian projection is simply a conditional spatial probability distribution that lets us know what mode M_B probabilistically "looks" like when a region p_i is active in cooccurring mode M_A . This is visualized in Figure 16.

We can equivalently define a Hebbian projection for a region $r \subseteq M_A$ constructed out of a subset of its codebook clusters $C_r = \{p_{r1}, p_{r2}, ..., p_{rk}\} \subseteq C_A$:

$$\vec{H}_{A}^{B}(r) = \left[\Pr(q_{1} \mid r), \Pr(q_{2} \mid r), ..., \Pr(q_{b} \mid r)\right]$$
(3.2)

We will also define the notion of a *reverse Hebbian projection*, which projects a Hebbian projection back onto its source mode. It lets us measure – from the perspective of

another modality – which other codebook regions in a slice appear similar to a reference region.



Figure 16 – A visualization of two Hebbian projections. On the top, we project from a cluster p_i in Mode A onto Mode B. The dotted lines correspond to Hebbian linkages and the blue shading in each cluster q_j in Mode B is proportional to $Pr(q_j|p_i)$. A Hebbian projection lets us know what Mode B probabilistically "looks" like when some prototype in Mode A is active. On the bottom, we see a projection from a cluster in Mode B onto Mode A.

To do this, we first define *weighted* versions of the functions defined above for a set of weights ω . Consider a region r, |r| = k, where each cluster is assigned some weight ω_i . We assume that $\sum \omega_i = 1$.

$$\begin{split} h_{\omega}(r) &= \sum_{p \in r} \omega_p h(p), \text{ where } p \text{ is a codebook cluster in region } r \\ \Pr_{\omega}(q,r) &= h_{\omega}(r,q) / h_{\omega}(r) = \sum_{p \in r} \omega_p h(p,q) / \sum_{p \in r} \omega_p h(p) \\ \vec{H}_{\omega}(r) &= \left[\Pr_{\omega}(q_1 \mid r), \Pr_{\omega}(q_2 \mid r), ..., \Pr_{\omega}(q_n \mid r) \right] \end{split}$$

The reverse Hebbian projection $\hat{H}_A^B(r)$ of a region $r \subseteq M_A$ onto mode M_B is then defined:

$$\hat{H}_{A}^{B}(r) = \vec{H}_{\vec{H}(r)}(M_{B})$$
(3.3)

$$= \left[\Pr_{H(r)}(p_1 | M_B), \Pr_{H(r)}(p_2 | M_B), ..., \Pr_{H(r)}(p_m | M_B) \right]$$
(3.4)

Again, when the modes are clear from context, we will simply refer to this as $\hat{H}(r)$.

This distribution has a simple interpretation: the reverse Hebbian projection from mode M_A onto mode M_B for some region $r \subseteq M_A$ is the Hebbian projection of all of mode M_B onto mode M_A , weighted by the forward Hebbian projection of region r, as shown in equation (3.3). This process is visualized in Figure 17. Note that we are projecting an entire mode M_B here. This might seem initially surprising, but it simply corresponds to a projection of a region that contains all the codebook clusters for a given slice.

The reverse Hebbian projection $\hat{H}(r)$ answers the question: what other regions does mode M_B think region r is similar to in mode M_A ? It can therefore be viewed as a distribution that measures cross-modal confusion. For this reason, it provides a useful optimization tool, because we will only need to disambiguate regions that appear in each other's reverse Hebbian projections, i.e., they have a non-zero (or above some threshold) probability of being confused for one another by other modalities.



Figure 17 – Visualizing a reverse Hebbian projection. We first generate the Hebbian projection of the green cluster p_i in Mode A onto Mode B. This projection is represented by the shading of each region q_j in Mode B, corresponding to $Pr(q_j|p_i)$. We then project all of Mode B back onto Mode A, weighting the contributions of each cluster q_i by $Pr(q_j|p_i)$. This generates the reverse Hebbian projection, which is indicated by the shading of regions in Mode A.

Measuring Distance in a Slice

Let us briefly review where we stand at this point. We have introduced the idea of a *slice*, which breaks up a representational space into many smaller pieces that are generated by hyperclustering it. We would like to assemble these small hyperclusters into larger regions that represent actual perceptual categories present in the input data. In this section, we define the non-Euclidean distance metric between the hyperclusters that helps make this possible.

Consider the colored regions in Figure 18. We would like to determine that the blue and red regions are part of their respective *blue* and *red* events, indicated by the colored ellipses. It is important to recall that the colors here are simply for the reader's benefit. There is no labeling of regions or perceptual events within the slice itself. We will proceed by formulating a distance metric that minimizes the distance between codebook regions that are actually within the same perceptual region and maximizes the distance



Figure 18 – Combining codebook regions to construct perceptual regions. We would like to determine that regions within each ellipse are all part of the same perceptual event. Here, for example, the two blue codebook regions (probabilistically) correspond the *blue* event and the red regions correspond to the *red* event.

between codebook regions that are in different regions. That this metric must be non-Euclidean is clear from looking at the figure. Each highlighted region is closer to one of a different color than it is to its matching partner.

We are going to use the Hebbian projections defined in the previous section to formulate this similarity metric for codebook regions. This will make the metric inherently crossmodal because we will rely on co-occurring modalities to determine how similar two regions within a slice are. Our approach is to compare codebook regions by comparing their Hebbian projections onto co-occurring slices. This process is illustrated in Figure 9.

The problem of measuring distances between prototypes is thereby transformed into a problem of measuring similarity between spatial probability distributions. The distributions are spatial because the codebook regions have definite locations within a slice, which are subspaces of \mathbb{R}^n . Hebbian projections are thus spatial distributions on *n*-dimensional data. It is therefore not possible to use one dimensional metrics, e.g., Kolmogorov-Smirnov distance, to compare them because doing so would throw away the essential spatial information within each slice.



Figure 19 – Our approach to computing distances cross-modally. To determine the distance between codebook regions r_1 and r_2 in Mode B on top, we project them onto a co-occurring modality (Mode A) as shown in the middle. We then ask: *how similar are their Hebbian projections onto Mode A?*, as shown on the bottom. We have thereby transformed a question about distance between regions into a question of similarity between the spatial probability distributions provided by their Hebbian projections.

Defining Similarity

What does it mean for two things to be similar? This deceptively difficult question is at the heart of mathematical clustering and perceptual categorization and is common to a number of fields, including computer vision, statistical physics, and information and probability theory. The goal of measuring similarity between different things is often cast as a problem of measuring distances between multidimensional distributions on descriptive features. For example, in computer vision, finding minimum matchings between image feature distributions is a common approach to object recognition (Belongie et al. 2002).

In this section, we present a new metric for measuring similarity between spatial probability distributions, i.e., distributions on multidimensional metric spaces. We will use this metric to compute distances between codebook regions by comparing their Hebbian projections onto co-occurring modalities, as shown above in Figure 19. Our approach is therefore inherently multimodal – although we may be unable to determine a priori how similar two codebook regions are in isolation (i.e., unimodally), we can measure their similarity by examining how they are viewed from the perspectives of other co-occurring sensory channels. We therefore want to formulate a similarity metric on Hebbian projections that tells us not how far apart they are but rather, how similar they are to one another. This will enable perceptual bootstrapping by allowing us to answer a fundamental question:

Can any other modality distinguish between two regions in the same codebook? If not, then they represent the same percept.

3.1.1 Introduction

There are a wide variety of metrics available to quantify distances between probability distributions (see the surveys in Rachev 1991, Gibbs and Su 2002). We may contrast these in many ways, including whether they are actually metrics (i.e., symmetric and satisfy the triangle inequality), the properties of their state spaces, their computational complexity, whether they admit practical bounding techniques, etc. For

example, the common χ^2 distance is not a metric because it is asymmetric. In contrast, Kolmogorov-Smirnov distance is a metric but is defined only over \mathbb{R}^1 . Choosing an appropriate metric for a given problem is a fundamental step towards solving it and can yield important insights into its internal structure.

In discussions of probability metrics, the notion of similarity generally follows directly from the definition of distance. Two distributions are deemed similar if the distance between them is small according to some metric; conversely, they are deemed dissimilar when the metric determines they are far apart. In our approach, we will reverse this dependency. We first intuitively describe our notion of similarity and then formulate a metric that computes it in a well-defined way. We call this metric the *Similarity distance* and it is the primary contribution of this section. Our approach is applicable to comparing distributions over any metric space and has a number of interesting properties, such as scale invariance, that make it additionally useful for work beyond the scope of this thesis.

3.1.2 Intuition

We begin by first examining similarity informally. Consider the two simple examples shown in Figure 20. Each shows two overlapping Gaussian distributions, whose similarity we would like to compare. Intuitively, we would say the distributions in Example A (on the left) are more similar to one another than those in Example B (on the right), because we will think of similarity as a measure of the overlap or proximity of spatial density. We are not yet concerned with formally defining similarity, but the intuition in these examples is exactly what we are trying to capture. Notice that the distributions in Example A cover roughly two orders of magnitude more area than those in Example B. Therefore, if we were to derive similarity from distance, the strong Euclidean bias incorporated into a wide variety of probability metrics would lead us to the opposite of our desired result. Namely, because the examples in B are much closer than those in A, we would therefore deem them more similar, thereby contradicting our desired meaning.



Figure 20 – Intuitively defining similarity. We consider the two distributions illustrated in Example A to be far more similar to one another than those in Example B, even though many metrics would deem them further apart due to inherent Euclidean biases. Notice that the distributions in Example A cover roughly two orders of magnitude more area than those in Example B. Note that simply normalizing the distributions before computing some metric on them would be ad hoc, very sensitive to outliers, and make common comparisons difficult.

Note that we cannot simply normalize pairs of distributions before computing some metric on them because our results would be highly sensitive to outliers. Doing so would also make common comparisons difficult, which is particularly important when demonstrating convergence in a sequence of probability measures. Finally, we want our similarity metric to be distribution-free and make no assumptions about the underlying data, which would make generalizing simple normalization schemes difficult.

3.1.3 Probabilistic Framework

We begin with some formal definitions. Our approach will be to define *Similarity distance* D_s as the ratio between two other metrics. These are the Kantorovich-Wasserstein distance and a new metric we introduce called the *one-to-many* distance. For each of these, we will provide a definition over continuous distributions and then present equivalent formulations for discrete weighted point sets. These are more computationally efficient for computing Similarity distance on the slice data structures introduced earlier. After this formal exposition, we intuitively explain and motivate these metrics in Section 3.6.4 and then show how Similarity distance is derived from them.

3.1.3.1 Kantorovich-Wasserstein Distance

Let μ and ν and be distributions on state space $\Omega = \mathbb{R}^n$. The Kantorovich-Wasserstein distance D_w (Kantorovich 1942, Gibbs and Su 2002) between μ and ν may be defined:

$$D_{W}(\mu,\nu) = \inf_{J} \{ D(x,y) : L(x) = \mu, L(y) = \nu \}$$
(3.5)

where the infimum is taken over all joint distributions J on x and y with marginals μ and ν respectively. For brevity, we will refer to D_w simply as the Wasserstein distance. Notice that in order to compute the Wasserstein distance, we already need to have a distance metric D defined to calculate the infimum. Where does D come from? In fact, in the approach described above, isn't D supposed to be the Similarity distance D_s , because we are proposing to use Similarity distance to measure distances within slices? Thus, we seem to have a chicken and egg problem from the start. We will sidestep this by defining D recursively through an iterative function system on D_s . This will allow us to compute Similarity distance by incrementally refining our calculation of it.

The definition in (3.5) assumes the distributions are continuous. Hebbian projections, however, are discrete distributions (i.e., weighted point sets) because they are over the codebooks within a slice. We may therefore simplify our computation by carrying it out directly over these codebooks. To do so, we define the Wasserstein distance on weighted point sets corresponding to discrete probability distributions. Consider finite sets $r_1, r_2 \subset \Omega$ with point densities φ_1, φ_2 respectively. Then we have:

$$D_{W}\left(\left\langle r_{1},\varphi_{1}\right\rangle,\left\langle r_{2},\varphi_{2}\right\rangle\right) = \inf_{J}\left\{D(x,y): L(x) = \left\langle r_{1},\varphi_{1}\right\rangle, L(y) = \left\langle r_{2},\varphi_{2}\right\rangle\right\}$$

which by (Levina and Bickel 2001) is equal to:

$$D_{W}\left(\langle r_{1}, \varphi_{1} \rangle, \langle r_{2}, \varphi_{2} \rangle\right) = \frac{1}{m} \min_{j_{1}, \dots, j_{m}} \sum_{i=1}^{m} \left[D\left(\langle r_{1}, \varphi_{1} \rangle_{i}, \langle r_{2}, \varphi_{2} \rangle_{j_{i}}\right)^{2} \right]^{1/2}$$
(3.6)

where *m* is the maximum of the sizes of r_1 and r_2 , the minimum is taken over all permutations of $\{1,...,m\}$, and $\langle r_a, \varphi_a \rangle_i$ is the *i*th element of set $\langle r_a, \varphi_a \rangle$. We note that

(ibid.) has shown this is equivalent to the *Earth Mover's distance* (Rubner et al. 1998), a popular empirical measure used primarily in the machine vision community, when they are both computed over probability distributions.

We can now define the Wasserstein distance between Hebbian projections of $r_1, r_2 \subseteq M_A$ onto M_B as:

$$D_{W}\left(\vec{H}(r_{1}), \vec{H}(r_{2})\right) = \inf_{J}\left\{D(x, y): L(x) = \vec{H}(r_{1}), L(y) = \vec{H}(r_{2})\right\}$$
(3.7)

where the infimum is taken over all joint distributions J on x and y with marginals $\vec{H}(r_1)$ and $\vec{H}(r_2)$. By (3.6), we have this is equal to:

$$D_{W}\left(\vec{H}(r_{1}),\vec{H}(r_{2})\right) = \frac{1}{m} \min_{j_{1},...,j_{m}} \sum_{i=1}^{m} \left[D\left(\vec{H}(r_{1})_{i},\vec{H}(r_{2})_{j_{i}}\right)^{2} \right]^{1/2}$$
(3.8)

where *m* is the number of codebook regions in M_B , the minimum is taken over all permutations of $\{1,...,m\}$, and $\vec{H}(r)_i = i^{th}$ component of $\vec{H}(r)$.

We note that the Wasserstein distance presented above is not a candidate for measuring similarity. In fact, referring back to Figure 20, the red and blue distributions in Example A here are further apart as determined by the Wasserstein distance than those in Example B, i.e., D_W (Example A) > D_W (Example B), which does not capture our intended meaning of similarity.

3.1.3.2 Computational Complexity

The optimization problem in (3.6) was first proposed by Monge (1781) and is known as the Transportation Problem. It involves combinatorial optimization because the minimum is taken over $O(2^m)$ different permutations and can be solved by Kuhn's Hungarian method (1955, see also Frank 2004). However, by treating it as a flow problem, we instead use the Transportation Simplex method introduced by Dantzig (1951) and subsequently enhanced upon by Munkres (1957), which has worst case exponential time but in practice is quite efficient (Klee and Minty 1972). To get some insight into the structure of this problem, we take a moment to examine the complexity of determining exact solutions to it according to (3.8). Although this is not necessary in practice, it is instructive to see how the choice of mixture distributions influences the complexity of the problem and the implications this has for selecting perceptual features. Notice that in the minimization in equation (3.8), the vast majority of permutations can be ignored because we only need examine regions that have non-zero probabilities in the Hebbian projections. In other words, we could choose to ignore any region $q_i \subseteq M_B$ where $\max(\vec{H}(r_1)_i, \vec{H}(r_2)_i) \leq \varepsilon$ for some small ε . A conservative approach would set $\varepsilon = 0$, however, one can certainly imagine using a slightly higher threshold to simultaneously reduce noise and computational complexity.

We may estimate the running time of calculating D_w exactly by asking how many nonzeros values we expect to find in the Hebbian projections onto mode M_B of two regions in mode M_A . Let us suppose that mode M_B actually has *d* events (of equal likelihood) distributed over *m* codebook regions. How many codebook regions are there within each event? If the events do not overlap, then we expect that each perceptual event is covered by *m/d* codebook regions, due to the density normalization performed during codebook generation. In this case, the minimization must be performed over $O(2^{1+m/d})$ permutations. Alternatively, it is possible for all of the sensory events to overlap, giving an upper bound, worst case of *m* regions per event and $O(2^m)$ running time. Thus, the running time is a function of the event mixture distributions as much as it is the number of codebooks. When Hebbian distributions are "localized," in the sense they are confined to subsets of the codebook regions, the worst-case running time is closer to $O(2^{1+m/d})$.

We can optimize the computation by taking advantage of the fact that the projections are over identical codebooks, i.e., their spatial distributions are over the same set of points generated by the hyperclustering of M_B . In the general statement of the Transportation Problem, this need not be the case. We can therefore reduce the number of codebook regions involved by removing the intersection of the projections from the calculation. Where they overlap, namely, the distribution described by a normalized $\min(\vec{H}(r_1), \vec{H}(r_2))$, we know the Wasserstein distance between them is 0. Therefore, let:

$$\vec{H}'(r_1) = \vec{H}(r_1) - \min\left(\vec{H}(r_1), \vec{H}(r_2)\right)$$
(3.9)

$$\vec{H}'(r_2) = \vec{H}(r_2) - \min\left(\vec{H}(r_1), \vec{H}(r_2)\right)$$
(3.10)

$$\Delta = 1 - \sum \min\left(\vec{H}(r_1), \vec{H}(r_2)\right) \tag{3.11}$$

We then have:

$$D_{W}\left(\vec{H}(r_{1}), \vec{H}(r_{2})\right) = \Delta D_{W}\left(\vec{H}'(r_{1}), \vec{H}'(r_{2})\right)$$
(3.12)

In other words, the Wasserstein distance computed over a common codebook (3.12) is equal to the distance computed on the distributions ((3.9) and (3.10)) over their non-intersecting mass (3.11). (Note that we must normalize (3.9) and (3.10) to insure they remain probability distributions, but the reader may assume this normalization step is always implied when necessary.) When the distributions overlap strongly, which we previously identified as the worst case scenario, we can typically use this optimization to cut the number of involved codebooks regions in half. When the distributions do not overlap, this optimization provides no benefit, but as we have already noted, this is a best case scenario and optimization is less necessary. As a further enhancement, we could also establish thresholds for Δ to avoid calculating D_w altogether. For example, in the case where their non-intersecting mass is extremely small, we might chose to define $D_w = 0$ or some other approximation.

In summary then, the computational complexity of exactly computing the Wasserstein distance very much rests on the selection of mixture distributions over which it is computed. These in turn depend upon the *feature selection* used in our perceptual algorithms, which directly determine the distributions of sensory data within a slice. We say that "good" features are ones that tend to restrict Hebbian projections to smaller subsets of slices and to reduce the amount of overlap among detectable perceptual events. (We suspect one can directly formulate a measure of the entropy in features based on

these criteria but have not done so here.) Empirically, "good" features for computing the Wasserstein distance tend to be similar to the ones we naturally select when creating artificial perceptual systems. "Bad" features provide little information because their values are difficult to separate, i.e., they have high entropy. Later in the thesis, we will draw biological evidence for these theses idea from (Ernst and Banks 2002).

3.1.3.3 The One-to-Many Distance

We now introduce a new distance metric called the *one-to-many* distance. Afterwards, we examine this metric intuitively and show how it naturally complements the Wasserstein distance. We will use these metrics together to formalize our intuitive notion of similarity.

Let *f* and *g* be the respective density functions of distributions μ and ν on state space $\Omega = \mathbb{R}^n$. Then the one-to-many distance (D_{OTM}) between μ and ν is:

$$\begin{aligned} D_{OTM}(\mu, \nu) &= \int_{\mu} f(x) \cdot D_{W}(x, \nu) \, dx \\ &= \int_{\mu} \int_{\nu} f(x) \cdot g(y) \cdot d(x, y) \, dx dy \\ &= \int_{\nu} g(y) \cdot D_{W}(\mu, y) \, dy = D_{OTM}(\nu, \mu) \end{aligned}$$

We define this over weighted pointed sets as:

$$D_{OTM}\left(\langle r_{1}, \varphi_{1} \rangle, \langle r_{2}, \varphi_{2} \rangle\right) = \sum_{p_{i} \in r_{1}} \varphi_{i} D_{W}\left(p_{i}, \langle r_{2}, \varphi_{2} \rangle\right)$$
$$= \sum_{p_{i} \in r_{1}} \varphi_{i} \sum_{p_{j} \in r_{2}} \varphi_{j} d(p_{i}, p_{j})$$
$$= \sum_{p_{j} \in r_{2}} \varphi_{j} \sum_{p_{i} \in r_{1}} \varphi_{i} d(p_{i}, p_{j})$$
$$= D_{OTM}\left(\langle r_{2}, \varphi_{2} \rangle, \langle r_{1}, \varphi_{1} \rangle\right)$$

The one-to-many distance is the weighted sum of the Wasserstein distances between each individual point within a distribution and the entirety of another distribution. It is

straightforward to directly calculate D_w between a point and a distribution and computing D_{OTM} requires $O(m^2)$ time, where *m* is the maximum number of weighted points in the distributions. Also, we note from these definitions that D_{OTM} is symmetric.

From this definition, we can now formulate D_{OTM} between Hebbian projections of $r_1, r_2 \subseteq M_A$ onto M_B :

$$D_{OTM}\left(\vec{H}(r_1), \vec{H}(r_2)\right) = \sum_{i \in \vec{H}(r_1)} \omega_i D_W\left(i, \vec{H}(r_2)\right)$$
(3.13)

$$= \sum_{i \in \hat{H}(r_1)} \omega_i \sum_{j \in \hat{H}(r_2)} \omega_j D(i, j)$$
(3.14)

where $\omega_i = \vec{H}(r_1)_i$ and $\omega_j = \vec{H}(r_2)_j$. Recall that $\vec{H}(r)_i = i^{th}$ component of $\vec{H}(r)$.

We see this is symmetric:

$$D_{OTM}\left(\vec{H}(r_1), \vec{H}(r_2)\right) = D_{OTM}\left(\vec{H}(r_2), \vec{H}(r_1)\right)$$

The one-to-many distance is a weighted sum of the Wasserstein distances between each individual region in a projection and the other projection in its entirety. The weights are taken directly from the original Hebbian projections. This is represented by the term $\omega_i D_w(i, \vec{H}(r_2))$ in (3.13). We note that the Wasserstein distance between a single region and a distribution is trivial to compute directly, as shown in (3.14), and it can be calculated in O(m) time, where *m* is the number of codebooks in Mode B. Therefore, D_{OTM} can be computed in $O(m^2)$ time.



Figure 21 – A simpler example. Mode B is the same as in previous examples but the events in Mode A have been separated so they do not overlap. The colored ellipses show how the external *red* and *blue* events probabilistically appear within each mode.



Figure 22 – Hebbian projections from regions in Mode B onto Mode A. Notice that the Hebbian projections have no overlap, which simplifies the visualization and discussion in the text. However, this does not affect the generality of the results.

3.1.4 Visualizing the Metrics: A Simpler Example

Consider the example in Figure 21. We have modified Mode A, on the bottom, so that its events no longer overlap. (Mode B on top remains unchanged.) This will simplify the presentation but does not affect the generality of the results presented here. As before, the two world events perceived by each mode are delineated with colored ellipses for the benefit of the reader but the modes themselves have no knowledge of them. The Hebbian projections from two codebook regions in Mode B are shown in Figure 22. We see this example was designed so that the projections have no overlap, making it easy to view them independently.

We now give an intuitive interpretation of the two distance metrics, D_W and D_{OTM} , based on the classic statement of the Transportation Problem (Monge 1781). This problem is more naturally viewed with discrete distributions, but the presentation generalizes readily to continuous distributions. Consider the Hebbian projections from our example in isolation, as show in Figure 23. On the left, $\vec{H}(r_1)$ is shown in red, and $\vec{H}(r_2)$ is shown in blue on the right. The shading within each Voronoi region is proportional to its weight (i.e., point density) within its respective distribution.

In the Transportation Problem, we imagine the red regions (on the left) need to deliver supplies to the blue regions (on the right). Each red region contains a mass of supplies proportional to its shading and each blue region is expecting a mass of supplies proportional to its shading. (We know that mass being "shipped" is equal to the mass being "received" because they are described by probability distributions.) The one-to-many distance is how much work would be necessary to deliver all the material from the red to blue regions, if each region had to independently deliver its mass proportionally to all regions in the other distribution. Work here is defined as $mass \times distance$.

The Wasserstein distance computes the minimum amount of work that would be necessary if the regions cooperate with one another. Namely, red regions could deliver material to nearby blue regions on behalf of other red regions, and blue regions could receive material from nearly red regions on behalf of other blue regions. Nonetheless, we maintain the restriction that each region has a maximum amount it can send or receive, corresponding to its point density. This is why the Wasserstein distance computes the solution to the Transportation Problem, which is directly concerned with this type of delivery optimization.

Thus, we may summarize that D_{OTM} computes an unoptimized Transportation Problem, where cooperation is forbidden, and that D_W computes the optimized Transportation Problem, where cooperation is required.



Figure 23 – Visualizing D_w and D_{OTM} through the Transportation Problem. We examine the Hebbian projections onto Mode A shown in Figure 22. $\vec{H}(r_1)$ is shown in red on the left and $\vec{H}(r_2)$ is shown in blue on the right. Each region is shaded according to its point density. In the Transportation Problem, we want to move the "mass" from one distribution onto the other. If we define $work = mass \times distance$, then D_{OTM} computes the work required if each codebook region must distribute its mass proportionally to all regions in the other distribution. D_w computes the work required if the regions cooperatively distribute their masses, to minimize the total amount of work required.

3.1.5 Defining Similarity

We are now in a position to formalize the intuitive notion of similarity presented above. We define a new metric called the *Similarity distance* (D_s) between continuous distributions μ and ν :

$$D_{S}(\mu,\nu) = \frac{D_{W}(\mu,\nu)}{D_{OTM}(\mu,\nu)}$$

and over weighted point sets:

$$D_{S}(\langle r_{1}, \varphi_{1} \rangle, \langle r_{2}, \varphi_{2} \rangle) = \frac{D_{W}(\langle r_{1}, \varphi_{1} \rangle, \langle r_{2}, \varphi_{2} \rangle)}{D_{OTM}(\langle r_{1}, \varphi_{1} \rangle, \langle r_{2}, \varphi_{2} \rangle)}$$

We thereby define the Similarity distance between Hebbian projections of $r_1, r_2 \subseteq M_A$ onto M_B :

$$D_{S}\left(\vec{H}(r_{1}), \vec{H}(r_{2})\right) = \frac{D_{W}\left(\vec{H}(r_{1}), \vec{H}(r_{2})\right)}{D_{OTM}\left(\vec{H}(r_{1}), \vec{H}(r_{2})\right)}$$
(3.15)

The Similarity distance is the ratio of the Wasserstein to the one-to-many distance. It measures the optimization gained when transferring the mass between two spatial probability distributions if cooperation is allowed. Intuitively, it normalizes the Wasserstein distance. It is scale invariant (see Figure 24) and captures our desired notion of similarity.

An important note to avoid confusion: Because D_s is a distance measure based on similarity – and not a similarity measure – *it is smaller for things that are more similar and larger for things that are less similar*. So, for any distribution v, $D_s(v,v)=0$, expressing the notion that anything is (extremely) similar to itself.

We briefly examine the behavior of D_s at and in between its limits. Let $\vec{H}(r_1)$ and $\vec{H}(r_2)$ be identical Hebbian projections separated by some distance Δ . Then we have the following properties:



Figure 24 – Examining Similarity distance. Comparing the distributions in the two examples, we have D_s (Example A) $\ll D_s$ (Example B), which captures our intuitive notion of similarity.

1) We see that as the distributions are increasingly separated, the optimization provided in the Wasserstein calculation disappears:

$$\lim_{\Delta \to \infty} D_W \left(\vec{H}(r_1), \vec{H}(r_2) \right) = D_{OTM} \left(\vec{H}(r_1), \vec{H}(r_2) \right)$$

and therefore:

$$\lim_{\Delta \to \infty} D_s \left(\vec{H}(r_1), \vec{H}(r_2) \right) = 1$$

2) As the distributions are brought closer together, the Wasserstein distance decreases much faster than the One-To-Many distance:

$$\frac{\partial}{\partial \Delta} D_{W} \left(\vec{H}(r_{1}), \vec{H}(r_{2}) \right) > \frac{\partial}{\partial \Delta} D_{OTM} \left(\vec{H}(r_{1}), \vec{H}(r_{2}) \right)$$

3) As they approach, it eventually dominates the calculation:

$$\lim_{\Delta \to 0} D_W(\vec{H}(r_1), \vec{H}(r_2)) = 0 \text{ and therefore, } \lim_{\Delta \to 0} D_S(\vec{H}(r_1), \vec{H}(r_2)) = 0$$

4) So, we see that $0 \le D_s(\vec{H}(r_1), \vec{H}(r_2)) \le 1$ and D_s varies non-linearly between these limits.

On the next two pages, we visualize the dependence of D_s on the distance Δ and the angle θ between pairs of samples drawn from different distributions. We examine samples drawn from Gaussian and Beta distributions in Figure 25 and Figure 26 respectively.



Figure 25 – Effects on D_s as functions of the distance D_e and angle θ between Gaussian distributions. As the distance D_e or angle between θ two Gaussian distributions decreases, we see how their corresponding similarity distance D_s decreases non-linearly in the graph on the bottom. D_e is shown in red and θ is shown in blue.



Figure 26-Effects on D_s as functions of the distance D_e and angle θ between Beta distributions. As the distance D_e or angle between θ the two Beta distributions decreases, we see how their corresponding similarity distance D_s decreases non-linearly in the graph on the bottom. D_e is shown in red and θ is shown in blue.

3.1.5.1 A Word about Generality

We can use Similarity distance to compare arbitrary discrete spatial probability distributions. Several such comparisons are illustrated in Figure 27. These examples are important, because our being able to compute Similarity distances between these distributions means that we will be able to perceptually ground events drawn from their mixtures. That all our examples have so far involved mixtures of Gaussians has simply been for convenience. We will demonstrate later in this chapter that we can separate events corresponding to a wide assortment of mixture distributions such as the ones shown here.



Figure 27 – Comparing spatial probability distributions. In each slice, the green and blue points represent samples drawn from equivalent but rotated 2-dimensional distributions. For each example, we identify the source distribution and the Similarity distance between the green and blue points. (A) A 2-D beta distributions with a, b = 4. Note the low density of points in the center of the distributions and the corresponding sizes of the codebook regions. D_S = 0.22. (B) A 2-D uniform distribution. D_S = 0.11. (C) A 2-D Gaussian distribution with σ = (.15, .04). D_S = 0.67. (D) A 2-D Poisson distribution, with λ = (50,30) and scaled by (.8, .3). D_S = .55.



Figure 28 – Some familiar non-parametric probability distributions within codebooks. We explore hand writing recognition using D_s in Appendix 2.

We can also apply the Similarity distance to non-parametric distributions. For example, consider the familiar distributions shown in Figure 28. In Appendix 2, we will investigate using D_s for handwriting recognition and examine some of the interesting properties of codebooks constructed on contours.

Defining the Distance Between Regions

We now use Similarity distance to define the *Cross-Modal distance* (D_{CM}) between two regions $r_1, r_2 \in M_A$ with respect to mode M_B :

$$D_{CM}(r_1, r_2) = \left[(1 - \lambda) \left[D_E(r_1, r_2) \right]^2 + \lambda \left[\sqrt{2} D_S(\vec{H}(r_1), \vec{H}(r_2)) \right]^2 \right]^{1/2}$$
(3.16)

$$= \left[(1 - \lambda) \left[D_E(r_1, r_2) \right]^2 + 2\lambda \left(\frac{D_W \left(\vec{H}(r_1), \vec{H}(r_2) \right)}{D_{OTM} \left(\vec{H}(r_1), \vec{H}(r_2) \right)} \right)^2 \right]^{1/2}$$
(3.17)

where D_E is Euclidean distance and λ is the relative importance of cross-modal to Euclidean distance. Thus the distance between two regions within a slice is defined to have some component $(1-\lambda)$ of their Euclidean distance and some component (λ) of the Similarity distance between their Hebbian projections. This is illustrated in Figure 29.

In almost all uses of cross-modal distance in this thesis, we set $\lambda = 1$ and ignore Euclidean distance entirely. However, in some applications, e.g., hand-writing or drawing recognition, spatial locality within a slice is important because it is a fundamental component of the phenomenon being recognized. If so, we can use a lower of λ . Determining the proper balance between Euclidean and Similarity distances is an empirical process for such applications.

So far, we have only considered two co-occurring modes simultaneously to keep the examples simple. However, it is straightforward to generalize the definition to incorporate additional modalities, and the calculation scales linearly with the number of modalities involved. To define the cross-modal distance (D_{CM}) between two regions $r_1, r_2 \in M_A$ with respect to a set of co-occurring modes $M_I \in M$, we define:

$$D_{CM}(r_1, r_2) = \left[(1 - \lambda_E) \left[D_E(r_1, r_2) \right]^2 + 2 \sum_{M_I \in \mathcal{M}} \lambda_I \left[D_S(\vec{H}_A^I(r_1), \vec{H}_A^I(r_2)) \right]^2 \right]^{1/2}$$
(3.18)

where the contributions of each mode M_I is weighted by λ_I and we set $\lambda_E + \sum \lambda_I = 1$. For guidance in setting the values of the λ_I , we can turn to (Ernst and Banks 2002), who found that in intersensory influence, people give preference to senses which minimize the variance in joint perceptual interpretations, confirming an earlier prediction by (Welch and Warren 1986) about sensory dominance during multimodal interactions. This lends credence to our hypothesis in section 3.6.3.2 regarding the computational value of entropy minimization in the selection of perceptual features. We reexamine these issues in the dynamic model presented in Chapter 4.



Figure 29 – Calculating the *cross-modal distance* between codebook regions in a slice. The distance is a function of their local Euclidean distance and the how similar they appear from the perspective of a co-occurring modality. To determine this for regions r_1 and r_2 in Mode B on top, we project them onto Mode A, as shown in the middle. We then compute the Similarity distance of their Hebbian projections, as shown on the bottom.

3.1.6 Defining a Mutually Iterative System

In this section, we show how to use the cross-modal distance function defined above to calculate the distances between regions within a slice. This statement may seem surprising. Why is any elaboration required to use D_{CM} , which we just defined? There are two remaining issues we must address:

- We have yet to specify the distance function D used to define the Wasserstein distance in equations (3.7) and (3.8), which was also "inherited" in our definition of the one-to-many distance in equation (3.14).
- 2) By defining distances cross-modally, we have created a mutually recursive system of functions. Consider any two regions r_1, r_2 in mode M_A . When we calculate $D_{CM}(r_1, r_2)$, we are relying on knowing the distances between regions within another mode M_B , which are used to calculate $D_s(\vec{H}(r_1), \vec{H}(r_2))$. However, the distances between regions in M_B are calculated exactly the same way but with respect to M_A . So, every time we calculate distances in a mode, we are implicitly changing the distances within every other mode that relies upon it. And of course, this means its own inter-region distances may change as a result! How do we account for this and how do we know such a system is stable?

We will approach both of these issues simultaneously. Suppose we parameterize the distance function D in all of our definitions:

$$D_{W}\left(\vec{H}(r_{1}),\vec{H}(r_{2}),D\right) = \frac{1}{m} \min_{j_{1},\dots,j_{m}} \sum_{i=1}^{m} \left[D\left(\vec{H}(r_{1})_{i},\vec{H}(r_{2})_{j_{i}}\right)^{2} \right]^{1/2}$$
(3.19)

$$D_{OTM}\left(\vec{H}(r_{1}), \vec{H}(r_{2}), D\right) = \sum_{i \in \vec{H}(r_{1})} \omega_{i} D_{W}\left(i, \vec{H}(r_{2}), D\right)$$
(3.20)

$$D_{S}\left(\vec{H}(r_{1}), \vec{H}(r_{2}), D\right) = \frac{D_{W}\left(\vec{H}(r_{1}), \vec{H}(r_{2}), D\right)}{D_{OTM}\left(\vec{H}(r_{1}), \vec{H}(r_{2}), D\right)}$$
(3.21)

$$D_{CM}(r_1, r_2, D) = \left[(1 - \lambda) \left[D_E(r_1, r_2) \right]^2 + 2\lambda \left[D_S(\vec{H}(r_1), \vec{H}(r_2), D) \right]^2 \right]^{1/2}$$
(3.22)

We now define an iterative function system on modes M_A and M_B that *mutually* calculates D_{CM} over their regions:

Let $\Delta_t^X = D_{CM}$ in mode M_X at time *t*. Recall that D_E is Euclidean distance. For all pairs of regions $r_i, r_j \in M_A$ and $q_i, q_j \in M_B$, we define:

$$\Delta_0^A(r_i, r_j) = D_E(r_i, r_j) \tag{3.23}$$

$$\Delta_0^B(q_i, q_j) = D_E(q_i, q_j)$$
(3.24)

$$\Delta_t^A(r_i, r_j) = D_{CM}\left(r_i, r_j, \Delta_{t-1}^B\right)$$
(3.25)

$$\Delta_t^B(q_i, q_j) = D_{CM}\left(r_i, r_j, \Delta_{t-1}^A\right)$$
(3.26)

Thus, we are start by assuming in (3.23) and (3.24) that the distances between regions in a slice are Euclidean, in the absence of any other information. (We later eliminate this assumption in the intermediate steps of cross-modal clustering, where we have good estimates on which to base the iteration.) The iterative steps are shown in (3.25) and (3.26), where at time *t*, we recalculate the distances within each slice based upon the distances in the other slice at time t-1. For example, notice how the definition of $\Delta_t^A(r_i, r_j)$ calculates D_{CM} using Δ_{t-1}^B in (3.25). After all pairs of distances have been computed at time *t*, we can then proceed to compute them for time t+1. As we did in equation (3.18), we can easily generalize this system to include any number of mutually recursive modalities. The complexity again scales linearly with the number of modalities involved. We stop the iteration when Δ_t^A and Δ_t^B begin to converge, which empirically tends to happen very quickly. Thus, we stop iterating on mode M_X at time *t* when:

$$\max_{r_i,r_j \in M_X} \frac{\left|\Delta_t^X(r_i,r_j) - \Delta_{t-1}^X(r_i,r_j)\right|}{\Delta_t^X(r_i,r_j)} < \kappa, \text{ for } \kappa = .9, \text{ we typically have } t \le 4.$$

We will refer to this final value of Δ_t^X for any regions $r_i, r_j \in M_X$ as $\tilde{D}_{CM}(r_i, r_j)$.

With this, we can complete our formal definition of the *slice* data structure. The final component necessary for specifying the topological manifold defined by a slice was the non-Euclidean distance metric between the hyperclustered regions. We now define this distance to be \tilde{D}_{CM} .

Cross-Modal Clustering

Recall that our goal has been to combine codebook regions to "reconstruct" the larger perceptual regions within a slice. The definition of the iterated cross-modal distance \tilde{D}_{CM} in the previous section allows us to proceed, because it suggests how to answer the following fundamental question:

Can any other modality distinguish between two regions in the same codebook? If not, then they represent the same percept.

Because \tilde{D}_{CM} represents the distance between two regions from the perspective of other modalities, we will use it to define a metric that determines whether to combine them or not. If $\tilde{D}_{CM}(r_i,r_j)$ is sufficiently small for two regions $r_1, r_2 \subseteq M_A$, then we will say they are *indistinguishable* and therefore part of the same perceptual event. If $\tilde{D}_{CM}(r_i,r_j)$ between two regions is large, we will say they are *distinguishable* and therefore, cannot be part of the same perceptual event. These criteria suggest the general structure of our cross-modal clustering algorithm. One important detail remains: how small must $\tilde{D}_{CM}(r_i,r_j)$ be for us to say it is "sufficiently" small? How do we define the threshold for merging two regions? An earlier version of this work appears in (Coen 2005).

3.1.7 Defining Self-Distance

We define the notion of *self-distance*, which measures the internal value of \tilde{D}_{CM} within an individual region. Thus, rather than measure the distance between two different regions, which has been our focus so far, *self-distance* measures the internal cross-modal distance between points within a single region. It is a measure of internal coherence and will allow us to determine whether two different regions are sufficiently similar to merge. We first define self-distance within an individual codebook cluster and then generalize this for regions composed of these clusters. Consider a slice M_A with associated codebook $C_A = \{p_1, p_2, ..., p_a\}$, generated from training dataset $T \subseteq \mathbb{R}^N$. Note the points in T represent the perceptual inputs to the slice that must be gathered before a codebook can be generated – it is not possible to hypercluster a slice that has no data within it. Let T_i be the Voronoi partitioning of Twith respect to the codebook clusters in C_A ; in other words, each T_i contains the points in the training dataset that are assigned to the cluster defined by p_i .

We are going to further partition each T_i into two sets, T_i^+ and T_i^- , by fitting a linear orthogonal regression onto it. For $T \subseteq \mathbb{R}^N$, this will generate an (N-1)-dimensional hyperplane that divides T_i into two sets, T_i^+ and T_i^- , minimizing the perpendicular distances from them to the hyperplane. Note that because the data are drawn from independent distributions, there is no error-free predictor dimension that generates the other dimensions according to some function. This is equivalent to the case where all variables are measured with error, and standard least squares techniques do not work in this circumstance. We therefore perform principal components analysis on T_i and generate the hyperplane by retaining its N-1 principal components. This computes what is known as the orthogonal regression and works even in cases where all the data in T_i are independent.

We use this hyperplane to partition T_i into T_i^+ and T_i^- , as shown in Figure 30. We define the *self-distance* (D_{self}) of codebook cluster p_i :

$$D_{self}(p_i) = \tilde{D}_{CM}\left(T_i^+, T_i^-\right)$$
(3.27)

The co-occurrence data for T_i^+ and T_i^- are collected simultaneously with that of their parent region p_i . They are gathered with respect to whole codebook regions in other slices — not with respect to those region's internal partitions. Therefore, there is minimal overhead in gathering Hebbian co-occurrence data for these subclusters and doing so



Figure 30 -- Visualizing the subpartitioning of codebook regions. The display on the right shows the internal partitioning of codebook regions in the slices on the left. Each cluster is divided into two components T_i^+ and T_i^- which are arbitrarily colored blue or green. The line that partitions each cluster is shown in red.

does not impact the computational complexity of this framework. The self-distance of a region is simply the weighted sum of the self-distances of its codebook clusters.

For a region $r = \{p_a, ..., p_b\}$, a subset of codebook clusters, let φ_i be the relative weight of each $p_i \in r$, e.g., $\varphi_i = |T_i| / \sum_{p_j \in r} |T_j|$. We define the *self-distance* of region r:

$$D_{self}(r) = \sum_{p_i \in r} \phi_i \ D_{self}(p_i)$$
(3.28)

3.1.8 A Cross-Modal Clustering Algorithm

We now present an algorithm for combining codebook clusters into regions that represent the sensory events within a slice. This is done in a greedy fashion, by combining the closest regions according to \tilde{D}_{CM} within each slice. We use the definition of *selfdistance* to derive a threshold for insuring regions are sufficiently close to merge. Afterwards, we examine the algorithm and some examples of its output.

Cross-Modal Clustering:

Given: A set of slices M and λ , the parameter for weighting Euclidean to Similarity distances. For each slice $M_i \in M$, we will call its codebook $C_i = \{p_1, ..., p_{k_i}\}$.

Initialization: For each slice $M_i \in M$, initialize a set of *regions* $R_i = C_i$. Each slice will begin with a set of regions based on its codebook. We will merge these regions together in the algorithm below.

Algorithm:

Calculate \tilde{D}_{CM} over the slices in set M.

While (true) do:

Calculate \tilde{D}_{CM} over the slices in set M. Use current \tilde{D}_{CM} as t=0 value For each slice $M_i \in M$: Sort the pairs of regions in M_i , $r_a, r_b \in R_i$, by $\tilde{D}_{CM}(r_a, r_b)$ For each pair $r_a, r_b \in R_i$, in sorted order: If $D_{self}(r_a) + D_{self}(r_b) > \tilde{D}_{CM}(r_a, r_b)$: Merge (r_a, r_b) Exit inner for loop. For each codebook cluster p_i in C_i : Let $r = \min_{r \in R_i} \arg[\tilde{D}_{CM}(p_i, r)]$ Move p_i into region rIf no regions were merged in any slice Either wait for new data or stop

 $\begin{aligned} r_a &= r_a \cup r_b. \\ R_i &= R_i / r_b. \\ \text{For all } p_i, p_j \in r_a, \text{ set } \tilde{D}_{CM} \left(p_i, p_j \right) = 0. \end{aligned}$

The cross-modal clustering algorithm initially creates a set of regions in each slice corresponding to its codebook. The goal is to merge these regions based on their cross-modal distances. The algorithm proceeds in a two-step greedy fashion:

- 1) For each slice, consider its regions in pairs, sorted by \tilde{D}_{CM} . If we find two regions satisfying $D_{self}(r_a) + D_{self}(r_b) > \tilde{D}_{CM}(r_a, r_b)$, we merge them and move onto the next step.
- 2) If as a result of this merger, some cluster p_i is now closer to another region, we simply move it there.

When we merge two regions, we set the pairwise distances between all codebooks clusters within them to 0, because we now view them as all part of the same underlying perceptual event and therefore equivalent to one another. At the end of each loop, we recompute \tilde{D}_{CM} using the current value as the starting point in the iteration, which propagates the effects of mergers to the other slices in M. In the event no mergers are made in any slices, we can choose to either wait for new data, which will update the Hebbian linkages, or we can terminate the algorithm, if we assume sufficient training data has already been collected.

Most clustering techniques work by iteratively refining a model subject to an optimization constraint. The iterative refinement in our algorithm occurs in the recalculation of \tilde{D}_{CM} , which is updated after each round of mergers within the slices. This spreads the effect of a merger within a slice by changing the Similarity distances between Hebbian projections onto it. This in turn changes the distances between regions in other slices, as discussed in section 3.7.1. The optimization constraint is that we do not merge regions whose internal self-distances sum to less than their \tilde{D}_{CM} distance. If we think of self-distance as measure of how much variation is permitted with an individual region, we only merge regions when the sum of their internal variations can account for their distance, as specified by the condition $D_{self}(r_a) + D_{self}(r_b) > \tilde{D}_{CM}(r_a, r_b)$.



Figure 31 – The progression of the cross-modal clustering algorithm. (A) shows the initial codebook creation in each slice. (B) and (C) show intermediate region formation. (D) shows the correctly clustered outputs, with the confusion region between the categories indicated by the yellow region in the center. Note in this example, we set $\lambda = .7$ to make region formation easier to see by favoring spatial locality. The final clustering was obtained by setting $\lambda = 1$.



Figure 32 – The output of cross-modally clustering four overlapping Gaussian distributions in each slice. The confusion region between them is indicated in the center of the clusters.



Figure 33 – Finding one cluster embedded in another. In mode B, cross-modal clustering is able both to detect the small cluster embedded in the larger one and to use this separation of clusters to detect those in mode A. This is due to the non-Euclidean scale invariance of Similarity distance, which is used for determining the cross-modal distance between regions. Thus, region size is unimportant in this framework, and "small" regions are as effective in disambiguating other modes as are "large" regions.



Figure 34 – Self-supervised acquisition of vowels (monophthongs) in American English. This work is the first unsupervised acquisition of human phonetic data of which we are aware. The identifying labels were manually added for reference and ellipses were fit onto the regions to aid visualization. All data have been normalized. Note the correspondence between this and the Peterson-Barney data show below.



Figure 35—The Peterson-Barney dataset. Note the correspondence between this and Figure 34. DRAFT NOTE: UPDATE THE COLORS AND IPA LABELS TO MATCH!!!

The progression of the algorithm starting with the initial codebook is shown in Figure 31. Setting $\lambda < 1$ includes Euclidean distances in the calculation of \tilde{D}_{CM} . This favors mergers between adjacent regions, which makes the algorithm easier to visualize. At the final step, setting $\lambda = 1$ and thereby ignoring Euclidean distance allows the remaining spatially disjoint regions to merge. Had we been uninterested in visualizing the intermediate clusterings, we would have set $\lambda = 1$ at the beginning. Doing so yields an identical result with this dataset, but the regions merge in a different, disjoint order. Note in general, however, it is not the case that different values of λ yield identical clusterings. All other examples in this thesis use $\lambda = 1$ exclusively.

Figure 32 demonstrates that the algorithm is able to resolve multiple overlapping clusters, in this case, two mixtures of four Gaussian distributions. Figure 33 show an important property of the Similarity distance, namely, it is scale invariant. The smaller cluster in Mode B is just as "distinct" as the larger one in which it is embedded. It is both detected and used to help cluster the regions in Mode A.

3.1.9 Clustering Phonetic Data

In Chapter 2, we asked the basic question of how categories are learned from unlabelled perceptual data. In this section, we provide an answer to this question using cross-modal clustering. We present a system that learns the number (and formant structure) of vowels (monophthongs) in American English, simply by watching and listening to someone speak and then cross-modally clustering the accumulated auditory and visual data. The system has no advance knowledge of these vowels and receives no information outside of its sensory channels. This work is the first unsupervised machine acquisition of phonetic structure of which we are aware.

For this experiment, data was gathered using the same pronunciation protocol employed by (Peterson and Barney 1952). Each vowel was spoken within the context of an English word beginning with [h] and ending with [d]; for example, /ae/ was pronounced in the context of "had." Each vowel was spoken by an adult female approximately 90-140 times. The speaker was videotaped and we note that during the recording session, a small number of extraneous comments were included and analyzed with the data. The auditory and video streams were then extracted and processed.

Formant analysis was done with the Praat system (Goedemans 2001, Boersma and Weenink 2005), using a 30ms FFT window and a 14th order LPC model. Lip contours were extracted using a system written by the author described in Chapter 2. Time-stamped formant and lip contour data were fed into *slices* in an implementation of the work in this thesis written by the author in Matlab and C. This implementation is able to visually animate many of the computational processes described here. This capability was used to generate most of the figures in this thesis, which represent actual system outputs.

Figure 34 shows the result of cross-modally clustering formant data with respect to lip contour data. Notice the close correspondence between the formant clusterings in Figures 22 and 23, which displays the Peterson-Barney dataset introduced earlier. We see the cross-modal clustering algorithm was able to derive the same clusters with the same spatial topology, without knowing either the number of clusters or their distributions.

The formant and lip slices are shown together in Figure 36, where the colors show region correspondences between the slices. This picture exactly captures what we mean by mutual bootstrapping. Initially, the slices "knew" nothing about the events they perceive. Cross-modal clustering lets them mutually structure their perceptual representations and thereby learn the event categories that generated their sensory inputs. The black lines in the figure connect neighboring regions within each slice and the red line connect corresponding regions in different slices. They show a graph view of the clustering within each slice and illustrate how a higher-dimensional manifold may be constructed out of lower-dimensional slices. This proposes an alternative view of the structures created by cross-modal clustering, which we hope to explore in future work.



Figure 36 – Mutual bootstrapping through cross-modal clustering. This displays the formant and lip slices together, where the colors show the region correspondences that are obtained from cross-modal clustering. Initially, the slices "knew" nothing about the events they perceive. Cross-modal clustering lets them mutually structure their perceptual representations and thereby learn the event categories that generated their sensory inputs.

Summary

This chapter introduced *slices*, a neurologically inspired data structure for representing sensory information. Slices partition perceptual spaces into codebooks and then reassemble them to construct clusters corresponding to the actual sensory events being perceived. To enable this, we defined a new metric for comparing spatial probability distributions called *Similarity distance*; this allows us to measure distances within slices through cross-modal Hebbian projections onto other slices. We then presented an algorithm for *cross-modal clustering*, which uses temporal correlations between slices to determine which hyperclusters within a slice correspond to the same sensory events. The cross-modal clustering algorithm does not presume that either the number of clusters in the data or their distributions is known beforehand. We also examined the outputs and behavior of this algorithm on simulated datasets and on real data gathered in

computational experiments. Finally, using cross-modal clustering, we have shown that sensory systems can be perceptually grounded by bootstrapping off each other.

DRAFT NOTE: To do for this chapter:

Replace Figures 22 - 24 with new versions

Include high-level Matlab code in an appendix, at least for cross-modal clustering and coherence determination.

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